# A Non-Parametric Surrogate-based Test of Significance for T-Wave Alternans Detection

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Abstract—We present a non-parametric adaptive surrogate test that allows for the differentiation of statistically significant T-Wave Alternans (TWA) from alternating patterns that can be solely explained by the statistics of noise. The proposed test is based on estimating the distribution of noise induced alternating patterns in a beat sequence from a set of surrogate data derived from repeated reshuffling of the original beat sequence. Thus, in assessing the significance of the observed alternating patterns in the data no assumptions are made about the underlying noise distribution. In addition, since the distribution of noise-induced alternans magnitudes is calculated separately for each sequence of beats within the analysis window, the method is robust to data non-stationarities in both noise and TWA. The proposed surrogate method for rejecting noise was compared to the standard noise rejection methods used with the Spectral Method (SM) and the Modified Moving Average (MMA) techniques. Using a previously described realistic multi-lead model of TWA, and real physiological noise, we demonstrate the proposed approach reduces false TWA detections, while maintaining a lower missed TWA detection compared with all the other methods tested.

A simple averaging-based TWA estimation algorithm was coupled with the surrogate significance testing and was evaluated on three public databases; the Normal Sinus Rhythm Database (NRSDB), the Chronic Heart Failure Database (CHFDB) and the Sudden Cardiac Death Database (SCDDB). Differences in TWA amplitudes between each database were evaluated at matched heart rate (HR) intervals from 40 to 120 beats per minute (BPM). Using the two-sample Kolmogorov-Smirnov test, we found that significant differences in TWA levels exist between each patient group at all decades of heart rates. The most marked difference was generally found at higher heart rates, and the new technique resulted in a larger margin of separability between patient populations than when the SM or MMA were applied to the same data.

Index Terms—ECG, Noise, Surrogate Analysis, T-Wave Alternans, TWA.

## I. INTRODUCTION

T-WAVE alternans (TWA), referring to beat-to-beat variability in the timing or shape of ST-T complex on the surface electrocardiogram, was first reported in 1908 by

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Hering [1]. Although the phenomenon is widely understood to be an important indicator of risk of sudden cardiac death (SCD), [2]-[4], until the 1980s TWA was believed to be rare. In 1981 Adam *et al.* first reported the existence of  $\mu$ V-level Twave alternans, too small in amplitude to be detected visually at standard electrocardiogram display scales [5]. Follow-up studies demonstrated that the absence of significant TWA in a patient with congestive heart failure, low ejection fraction, or a recent myocardial infarction is strongly predictive of a low risk of SCD [6], [7]. A positive finding in such a patient, though less specific, may indicate that an implantable cardiac defibrillator would be appropriate, an indication that can be confirmed using invasive testing. However, the positive predictive value of TWA remains low [8], and it is yet to be determined whether further improvements in the methodology of TWA detection/quantification can improve the positive diagnostic power of the TWA test.

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One unresolved issue in the area of TWA analysis is that of noise modeling and rejection of false detections while maintaining a low level of missed detections [9]–[11]. A comprehensive list of various TWA estimation and detection techniques is provided in Martínez *et al.* (2005) [12]. Two of the most common shortcomings of the discussed methods of TWA detection are 1) unjustified assumptions about the nature of the physiological noise (*e.g.*, Gaussian or Laplacian distributions) [11], and 2) arbitrary detection thresholds, often tuned on patient populations that are judged as healthy [13], [14].

In this article, we seek to determine if in the presence of noise (due to exogenous sources such as electrode movements or endogenous interferences such as muscle artifacts) alternating-like patterns can appear in the data, and whether in the absence of an appropriate statistical test such patterns can be mistaken for physiological-based TWA. We propose a non-parametric test to mitigate the problem of TWA false detection. The proposed test makes no assumption concerning the distribution or stationarity of the noise or the TWA in the data, and therefore is robust under varying recording conditions. The purpose of this work is to devise a robust statistical test to assist in accurate detection of TWA, independent of the particular estimation algorithm being used. To the best of our knowledge, this work is the first to propose a statistical test for TWA detection that is completely non-parametric and makes no assumption about the nature (distribution or dynamics) of the underlying noise or the TWA activity itself.

To provide a comparative study of the proposed TWA detection algorithm, we used an open source TWA analysis

tool to evaluate current standards for TWA metrics on four datasets. First, by using a model of TWA, to which realistic noise is added, we created a gold standard dataset in which the existence and magnitude of TWA is completely known. We then evaluated the concept of false estimation of TWA amplitude by the standard TWA analyzers at low levels of TWA amplitude and varying noise level to determine the sensitivity floor of various noise rejection techniques. Once the range of the standard TWA analysis techniques was determined, we investigated the feasibility of assessing statistical significance of a given alternans amplitude via a non-parametric surrogate test that allows for the differentiation of statistically significant T-Wave Alternans from alternating patterns that can be solely explained by the statistics of noise. Our surrogate method is similar to that described by Small et al. [15] and Theiler [16], [17], except that all the computations are performed in the time domain rather than the frequency domain. We also note that our approach is related to the approximate permutation test, Monte Carlo permutation tests or random permutation tests [18]. The proposed statistical significance test was then applied to three publicly available databases to investigate reports that TWA manifest more significantly at higher heart rates in both normal and cardiac-impaired populations [19].

We start with a brief description of the datasets utilized in this work, followed by an introduction to the most commonly utilized TWA analysis methods and a discussion of false detections in the presence of noise. Next, we describe the proposed non-parametric statistical test to separate real TWA effect from the noise induced alternans-like artifacts. Finally, we evaluate the performance of the proposed approach compared to the standard approaches using simulated vectorcardiographs (VCGs) with known TWA amplitude and three publicly available databases.

#### II. MATERIALS AND METHODS

Four data sets were used for the analysis; one set of computer simulated VCGs with known TWA amplitude and additive physiological noise (from the MIT-BIH Noise Stress Test Database), one set of recordings from healthy subjects, one set with chronic heart failure, and one set of recordings from sudden cardiac death patients.

## A. Simulated TWA

Five minutes duration records with TWA amplitudes of 0 through 100  $\mu$ V were generated using an artificial multilead VCG model with realistic TWA-like effects [20] and the x-axis of the VCG was chosen as the test signal. Next, noise segments of five minutes duration with random starting points were selected from the MIT-BIH Noise Stress Test Database (NSTDB) [21]. The NSTDB comprises recordings of three different types of noise, namely baseline wander, electrode movement, and muscle artifacts. The additive noise was constructed by mixing all three noise types and the power of the noise with respect to the VCG signal was adjusted to simulate records of SNR of 10, 20, 30 dB, and no noise (due to space limitations only results from SNR of 10 dB and no noise are reported). For a given SNR level and TWA amplitude, we generated 50 VCG records of five minutes each at a sampling



Fig. 1. Examples of simulated VCG with TWA amplitude of 23  $\mu$  V. Physiological noise consisting of a mixture of muscle artifacts, electrode movements, and baseline wander are added to each record. Only simulations at SNR of 10 dB (top) and clean VCG (bottom) are shown here. Zooming into the bottom plot one can observe the micro-volt variations from a normal beat (type-A) to an abnormal beat (type-B). The maximum amplitude variation between a type-A and a type-B beat is concentrated around the T-wave peak.

frequency of 500Hz and 16-bit resolution per sample, which is sufficient to prevent significant quantization noise [10], [22]. Individual records differ in that 1) the underlying VCGs were generated using a stochastic model of heart rate variability (70±5 beats/min) [23] and 2) the additive noise was taken from random five minute segments of the NSTDB. A short segment from one of the simulated records with a TWA amplitude of 23  $\mu$ V with additive noise at 10 dB SNR and no noise is shown in Fig. 1.

#### B. Real ECG Recordings

To assess the effectiveness of the proposed statistical test, we compared the performance of each of the described TWA detection methods for separating patient populations according to the magnitude of TWA activity they manifest. To this end we employed three publicly available databases to investigate reports that TWA manifest more significantly at higher heart rates, and more often within the cardiac-impaired populations [19].

#### 1) Normal Sinus Rhythm Database (NSRDB):

This database includes 18 long-term (at least 8 hour long) ECG recordings of subjects referred to the Arrhythmia Laboratory at Boston's Beth Israel Deaconess Medical Center. Subjects included in this database were found to have had no significant arrhythmias; they include 5 men, aged 26 to 45, and 13 women, aged 20 to 50 years. Recordings were performed at 128Hz sampling frequency and 12-bit resolution [24].

2) Chronic Heart Failure Database (CHFDB):

This database includes long-term ECG recordings from 15 subjects (11 men, aged 22 to 71, and 4 women, aged 54 to 63 years) with severe congestive heart failure. This group of subjects was part of a larger study group receiving conventional medical therapy prior to receiving the oral inotropic agent,

milrinone. Recordings were performed at 250Hz sampling frequency and 12-bit resolution [24].

## 3) Sudden Cardiac Death Database (SCDDB):

These data include 23 patients with underlying sinus rhythm (four with intermittent pacing), one who was continuously paced, and four with atrial fibrillation. All patients had a sustained ventricular tachyarrhythmia, and most had an actual cardiac arrest. The recordings were performed at sampling frequency of 250Hz and 12-bit resolution [24].

#### C. TWA Estimation and Detection

Since the purpose of this study was to propose a robust test of significance of TWA patterns, independent of the particular preprocessing (that is, pre-filtering, QRS detection and beat alignment) or estimation method, we utilized the same preprocessing steps across all methods. (For a more thorough description see [25].) Our implementations of the MMA and SM are based on descriptions found in Martínez *et al.* (2005) [12] and Narayan (2006) [26] and are described in the following sections.

The algorithms and metrics chosen for comparative study in this work were intended to mimic the approaches employed in commercial equipment and used most often by clinicians rather than to provide an exhaustive comparison of all TWA analysis techniques. To facilitate comparisons across various methods an analysis window of length L=64 beats with 32 beats overlap was utilized independent of the particular TWA algorithm. All the analyses in this article were performed on a single lead of the ECG records (lead I). Although different subjects may manifest maximal TWA activity across different leads, we expect the differences to average out over our databases.

## 1) The Proposed Detection Method (Surrogate Data Analysis):

In this work, we propose a non-parametric (assumption-free) statistical test to separate physiologically-induced alternating patterns (real TWA) in a beat sequence from that which could be a byproduct of the way one measures TWA amplitude and deals with the artifacts of recording noise. The main motivation behind the surrogate data analysis (SDA) method is that if the estimated alternans amplitudes are not artifacts of noise, then by eliminating the temporal relationship between the beats- through shuffling of the beat sequence- the amplitude of the beat-to-beat alternation ought to decrease significantly. Henceforth, we define a noise induced alternating pattern (NIAP) as an alternating pattern in a beat sequence that is caused by factors other than alternation in ventricular repolarization on an every-other-beat basis.

To cast the problem into a more rigorous statistical framework, one has to approximate the distribution of NIAP. A surrogate measure of NIAP may be obtained through repeated reshuffling of the beat sequence (say N = 250 times) and estimating the alternans amplitude for each surrogate arrangement of beats. In general, as the number of surrogates (shufflings) increases, the normalized histogram of the measured NIAP will approach the true distribution of NIAP. A statistical test can then be constructed by comparing the measured TWA amplitude against some upper percentile ( $(1 - \alpha) \times 100$ ) of the NIAP estimates (e.g.,  $95^{th}$  percentile or  $99^{th}$  percentile for  $\alpha = 0.05$  or  $\alpha = 0.01$ , respectively). If the estimated TWA amplitude is greater than or equal to all the NIAP values up to and including the  $(1-\alpha) \times 100$  percentile, the estimated TWA amplitude is significant and its value is reported. Otherwise, the TWA amplitude is labeled *indeterminate* for the given analysis window. Thus, the indeterminate cases are those which neither the presence nor the absence of TWA activity can be ruled out. (It is worth noting that, in the absence of any TWA activity and no random beat-to-beat variations, all possible random arrangements of beats must result in 0 V TWA amplitude, and thus alternans-free beat sequences would not be labeled indeterminate. However, in real data due to the presence of noise certain arrangements of beats will result in non-zero TWA amplitude, therefore, as a consequence of our definition of indeterminacy and noise, 0 V TWA amplitude is almost always labeled indeterminate. However, this does not cause any problem since in practice, missed or indeterminate 0 V alternans are unimportant).

Fig. 2 illustrates the normalized histogram of the NIAP (black) calculated (using the simple averaging method described in the following section) from a 64 beats long segment of the simulated ECG shown in the upper panel of Fig. 1. Superimposed on the graph are the fitted gamma distribution (dark blue) calculated from 250 times reshuffling of the beats and the  $95^{th}$  and  $99^{th}$  percentiles of the calculated alternans amplitude (of the empirical distribution) and fitted (parametric) gamma distribution. If the estimated TWA amplitude of the unshuffled beat sequence is larger than the  $95^{th}$  or  $99^{th}$ percentiles of the NIAP distribution one can confidently reject the null hypothesis (*i.e.*, the alternating pattern in the beat sequence can be explained by the statistics of noise) at  $\alpha = 0.05$  or  $\alpha = 0.01$ , respectively. Note that, by introducing a parametric form one can incorporate a belief pertaining to the tail of the distribution or frequency of rare events (heavytailed vs. light-tailed) which may not be captured through a moderate number of reshufflings. For instance, a heavy-tailed distribution can further reduce false alarm rates, since the upper percentiles of such distribution will be further to the right of the corresponding percentiles of the empirical distribution (or normalized histogram). However, this reduction in false alarm rates comes at a cost of increasing missed detections. This may be a large contributing factor in reports that current TWA analysis approaches are specific but not sensitive.

In Fig. 3 the 99<sup>th</sup> percentile of NIAP amplitude in reshuffled beat sequences are shown for simulated VCG records at SNR of 10 dB and no noise. Each red square on the graph represents the median over 500 values (50 records of the same TWA amplitude and 10 overlapping windows per record). Within each analysis window (of length L = 64 beats) the beat sequence is reshuffled 250 times, the alternans amplitude is calculated for each unique arrangement of the beats, and the 99<sup>th</sup> percentile of alternans amplitude over all 250 arrangements is recorded. Noteworthy is the tendency of the 99<sup>th</sup> percentile to increase with the simulated TWA amplitude. Note also that the NIAP is non-zero (even for 0  $\mu V$  TWA) and that the baseline NIAP increases as the SNR drops. These



Fig. 2. Normalized histogram of the NIAP (black) and fitted gamma distribution (dark blue) calculated from 250 times reshuffling of the beats within an analysis window (L = 64 beats). Marked are the 95<sup>th</sup> and 99<sup>th</sup> percentiles of the calculated alternans amplitude (non-parametric) and the fitted gamma distribution (parametric). If the estimated TWA is larger than the 95<sup>th</sup> or 99<sup>th</sup> percentile one can confidently reject the null hypothesis at  $\alpha = 0.05$  or  $\alpha = 0.01$ , respectively.

observations can be explained by the fact that after reshuffling, the number of type-A and type-B beats within the even and the odd group of beats is equal, then any difference between the average value of even group and odd group will be due to noise, since the type-A beats (type-B beats) within the even group will cancel the type-A beats (type-B beats) within the odd group. When the reshuffling of the beat sequence is thoroughly random, certain arrangements of beats may result in one of the beat types being overly represented in the odd or even group of beats, and therefore, some of the inter-group differences will be due to the existence of distinct beat types rather than being purely a noise artifact. In any event, the point of reshuffling the beat sequence is that, if there are two distinct beat types that manifest themselves in an alternating scheme (ie.,  $ABABAB \cdots$ ) then almost all other arrangements of the beats ought to produce an equal or smaller average difference between the odd beats and the even beats.

It should be noted that there are L! ways to arrange L beats, and  $(L/2)! \times (L/2)!$  ways to arrange these beats such that the new arrangements result in the same set of even and odd group of beats as in the original beat sequence. In general, the latter number is negligibly smaller than the former, and thus, the probability of generating beat sequences with even and odd groups of beats similar to the original beat sequence is negligibly small (for 250 shuffles and L = 64, this probability is approximately:  $250 \times 6.9 \times 10^{70}/1.3 \times 10^{89} \approx 1.4 \times 10^{-16}$ ). Furthermore, even if by chance shuffling results in such an event, the associated alternans amplitude will belong to the tail of the NIAP distribution and hence will not cause a missed detection for even a conservative significance level of  $\alpha =$ 0.01.

In this work the SDA-based detection technique employs a simple averaging-based method for estimating TWA ampli-



Fig. 3. Range of NIAP amplitudes at  $95^{th}$  percentile significance in reshuffled beat sequences using the SAM (see Fig. 2), representing a statistical measure of the upper-limit on the NIAP. Median (boxes), 5% (lower line) and 95% (upper line) are plotted to illustrate spread of the  $99^{th}$  percentile at each simulated TWA amplitude and across all simulated records, at *SNR* of 10 dB and no noise scenario.

tude, which we now describe.

## 2) Simple Averaging Method:

The simple averaging (SAM) method is based on calculating the absolute value of the difference between the average of the even and odd groups of beats within the analysis window, at every sample point within the ST-T complex, and taking the maximum value of the calculated differences within the ST-T complex. The SAM method is only an amplitude estimation technique and is essentially equivalent to the amplitude estimation part of the spectral method with a rectangular window.

#### 3) Spectral Method:

In our implementation of the spectral method (SM) [5] [27], we utilized Welch's non-overlapping periodogram method of estimating power spectral density and a Hamming window [12]. The alternans value was considered significant if the kvalue was larger than 3 (where the k-value refers to the spectral ratio index utilized within the SM method for detection purposes [27]. It can be shown that the periodogram calculated at the frequency of 0.5 cycles/beat is proportional to calculating the difference of even and odd group of beats (or a windowed version of the even and odd beat sequences in the case of a Hamming window) [12], [28].

## 4) Modified Moving Average Method:

The modified moving average (MMA) method was devised as an ad hoc method of calculating average templates for the even and the odd group of beats that are less sensitive to large fluctuations in T-Wave amplitude [29]. The reported TWA amplitude is the maximum value of the difference between the calculated even and odd templates. Note also that, the MMA method is only a TWA amplitude estimation technique and does not include any test of significance (that is, it performs TWA amplitude estimation but no TWA detection). In practice, certain steps in the preprocessing/alignment stepsuch as exclusion of beats with abnormal fluctuations in the TP segment (from the end of T-wave to beginning of the Pwave)- are taken to reduce noise artifacts [30]. However, since in this work a uniform preprocessing/alignment step is utilized independent of the particular TWA detector, the TP segment based noise rejection was omitted.

#### **III. RESULTS**

#### A. Simulated Data

In this section we compare the performance of the SDA, SAM, SM, and MMA methods on the simulated data described in section II-A, with and without additive realistic noise. The MMA method, as noted in section II-C4, is only a metric of TWA amplitude and does not include an explicit detection step. In contrast, SM uses thresholding method to reject noise artifacts. Furthermore, the SAM method is employed with and without our proposed significance testing in order to establish a baseline performance. Note that, the significance level for the surrogate test was picked to yield the same level of specificity (*i.e.*, proportion of negatives which are correctly identified) at  $0\mu$ V TWA amplitude as the SM method with a k-value of 3. In general a more stringent level of specificity would necessarily result in a larger percentage of missed detections and vice versa. Thus, by fixing the level of specificity of all detection algorithms one can compare their missed detection rates, as a means to assess their relative performance.

#### 1) Performance of the SDA and SAM:

Fig. 4 (top) illustrates the performance of the SAM without significance testing on the simulated data at SNR of 10 dB and no noise scenario. Represented in each figure are the lower 5%, median, and the upper 95% of the estimated TWA amplitude, as well as the identity line y = x (representing ideal estimation). At each given SNR there was a noise floor that hindered accurate detection of TWAs with small amplitudes. This noise floor decreased with an increase in SNR and resulted in false quantification of TWA amplitude, particularly at low TWA amplitudes (note that, even though only the results for SNR of 10 dB is shown here, these observations were consistent for SNRs of 20 and 30 dB). Even in the absence of background noise, a lower noise floor of 5-10  $\mu$ V was found below which it was impossible to distinguish real TWA from noise artifacts. Note that, in Fig. 4 (top) since the estimated alternans were all accepted (in the absence of significance testing) the percentage of indeterminate cases (or missed detections) were zero in all cases.

Fig. 4 (bottom) presents the TWA detection statistics after rejecting cases that were ruled false positives using the SDA method ( $\alpha = 0.01$ ), as well as, the percentage of indeterminate cases (grey color error bar). Note also that at the SNR of 10 dB and TWA amplitude of  $0\mu$ V approximately 99% of the NIAP were rejected. Furthermore, as we show next, the percentage of missed detections at higher TWA amplitudes were notably smaller than the SM (*e.g.*,  $20 \pm 15\%$  at the largest simulated TWA amplitude and SNR of 10 dB, as apposed to  $50 \pm 20\%$ for the SM).

#### 2) Performance of the SM:

Performance of the spectral method on the equivalent data is presented in Fig. 5 (bottom), using a k-value of 3 that is assumed to be constant throughout the analysis. Note that the SM yielded a substantial rejection rate even at higher TWA amplitudes (e.g.,  $50 \pm 20\%$  at the largest simulated TWA amplitude and SNR of 10). Note also that, the SM underestimated the TWA amplitude by a constant factor.

3) Performance of the MMA method:



Fig. 4. Performance of the SAM on the simulated data before (top) and after (bottom) application of the SDA method ( $\alpha = 0.01$ ). Estimated TWA amplitude (Calc. TWA Amp.) versus simulated TWA amplitude at SNR of 10 dB (left) and noise free simulated VCGs (right) are shown. Each point on the figure is calculated from 50 simulated VCG records of 5 minutes length each. At a heart rate of  $70 \pm 5$  beats/min this results in roughly 10 TWA amplitude measurements per record, and thus a total of 500 estimates. Represented in each figure are the lower 5%, median, and the upper 95% of the estimated TWA amplitude, as well as, the line y = x (representing ideal detection). The grey color error bars represent the percentage of indeterminate cases (%indet.). Note that, at the SNR of 10 dB application of surrogate testing resulted in rejection of approximately 99% of episodes around  $0\mu V$  simulated TWA amplitude, and simultaneously the percentage of missed detections (or indeterminate cases) at higher TWA amplitudes is notably smaller than the SM method (see Fig. 5).

Fig. 5 (top) illustrates the performance of the MMA method on the simulated ECG. Comparing with the Fig. 4 (top) it can be seen that the modified averaging method employed by the MMA method tended to amplify the noise. For instance at the SNR of 10 dB, and in the absence of TWA the median of the MMA estimates was 40  $\mu$ V compared with 30  $\mu$ V in the case of the SAM. This observation affirms and complements the observations made by Cox *et al.* [31] who conclude that: "MMA amplifies TWA compared to traditional spectral analysis, but both likely reflect similar pathophysiology". However, our simulations indicate that MMA amplifies both TWA as well as the effect of the recording noise. Due to the nonlinear nature of the MMA method we were not able to single out a unique cause for this behavior.

Table I summarizes the performance of the two TWA detection algorithms discussed in this work.

#### 4) Effect of Window Size:

Although we fixed the number of beats in our analysis to 64 beats, to allow direct comparison between different detection techniques for noise rejection, different studies have employed varying analysis window lengths, ranging from 16 beats [30] for the MMA method and 128 beats for the SM [27]. To determine the influence of the analysis window length on the results reported in this work, we repeated all the simulation studies with a 32 beats and a 128 beats window. Decreasing the analysis window length led to a raising of the noise floor and an increase in the percentage of indeterminate cases, since

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Fig. 5. Performance of the MMA method (top) and the SM (bottom) using a k threshold value of 3. Note that, such choice of k results in rejection of almost  $98 \pm 2\%$  of false estimates at 0  $\mu$ V TWA amplitudes and rejection of  $50 \pm 20\%$  of estimated values at the largest simulated TWA amplitude and SNR of 10. See Fig. 4 for explanation of legend.

#### TABLE I

Performance summery of the three TWA detection algorithms discussed in this work at low SNR of 10dB, based on Figs. 4 (bottom,left) and 5 (bottom,left). In the case of the SM, the spectral ratio index k was set equal to 3. The significance level of the SDA method ( $\alpha = 0.01$ ) was chosen to yield similar false alarm rate at 0  $\mu$ V as the SM with k = 3 (*i.e.*,  $2 \pm 2\%$ ). Application of the Wilcoxon rank sum test indicates significance differences (indicated by †) between the two algorithms in terms of percentage of missed detections, at every simulated TWA amplitude level from 50 to  $100\mu$ V (p < 0.001). (Due to space limitations only results for  $50\mu$ V and  $100\mu$ V are presented here.)

	~	~	
Detection Method	% False Alarm	% missed detections	
	at 0 μV,	at 50 $\mu$ V 100 $\mu$ V,	
	SNR=10 dB	SNR=10 dB	
SM (spectral ratio test)	$2\%\pm2\%$	$80\% \pm 13\% \dagger 50\% \pm 20\% \dagger$	
SDA	$1\%\pm1\%$	$60\% \pm 25\% \dagger 20\% \pm 15\% \dagger$	

the noise reduction effect of averaging was less marked when using fewer beats, but did not influence the trend observed in our results. On the other hand, increasing the number of beats in a qualitatively similar way decreased the NIAP level across all the methods.

### B. Real ECG Recordings

In this section we present results of a comparative TWA analysis of three publicly available databases using the SAM and the SDA methods. Our goal was to investigate the effects of the proposed statistical test to facilitate analysis of the data (in terms of separability of patient populations according to their level of TWA activity) independent of the particular method of estimation of the TWA amplitude. To this end, we applied the SAM (the simplest estimation method) without and with significance testing to the three databases. To facilitate comparison, we also present the performance of the SM and the MMA method on the same databases.

Fig. 6 compares the effects of significance testing on three different patient populations. The top panels show the esti-



Fig. 6. Comparison of the NSRDB, CHFDB, and SCDDB patient populations (at matched HRs) using the SAM (top) and the SDA method (bottom). The small numbers by the open blue circles indicate the number of detected episodes of TWA for the given HR range. The grey error bars signify the percentage of indeterminate cases at each HR range over the entire population. Note that, in the top panels the indeterminate cases are caused by preprocessing failure of associated analysis windows, while in the indeterminate cases in the bottom panels are an aggregate result of preprocessing failure and application of the SDA method ( $\alpha = 0.05$ ). In comparison to the top panels, the number of detected episodes of TWA in the bottom panels are greatly reduced (see numbers by the open blue circles), and the margin of separability among patient populations is increased (see Table II).

mated TWA amplitude on the NSRDB, CHFDB, and SCDDB patient populations at matched heart rate (HR) decades using the SAM with no surrogate testing. We chose to break down the data into HR decades because TWA is hypothesized to be a HR dependent phenomenon [19]. By doing so, we avoid any bias due to the expected differences in HRs between each population or for any differences in the noise rejection abilities of each TWA method which may be HR-dependent. Note that, in the case of real data the definition of *indeterminate* is further extended to include preprocessing failure (due to misalignments, excessive ectopic beats within the analysis window, etc). The bottom panels in Fig. 6 present results of applying the proposed SDA method. Setting  $\alpha = 0.05$  resulted in rejection of a large number of alternans-like episodes which did not pass the test of significance.

Table II summarizes the differences between the NSRDB, CHFDB, and SCDDB populations, as depicted in Fig. 6. For the purpose of comparison, Tables IV and III summarize the performance of the MMA method and the SM to the same databases.

These figures and tables demonstrate that the SDA method is effective in separating the three patient populations according to the median of their TWA activity. For instance, before significance testing the difference of median TWA amplitude between the NSRDB and SCDDB populations for HR band of 110 - 120 was  $-1.35\mu$ V (see the  $4^{th}$  column of Table *III*). However, after removing episodes of false positive- using the SDA method- this difference was  $41.51\mu$ V, indicating a much higher level of TWA activity among the SCDDB patient population.

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#### TABLE II

COMPARISON OF TWA ACTIVITY AT DIFFERENT HR IN THE NSRDB, CHFDB, AND SCDDB POPULATIONS USING THE SAM WITHOUT SIGNIFICANCE TESTING (TOP) AND AFTER SIGNIFICANCE TESTING WITH  $\alpha=0.05$  (bottom). For a given HR range,  $\Delta_{med}(1,2)$  is the median TWA amplitude of CHFDB population minus the median TWA AMPLITUDE IN THE NSRDB POPULATION. SIMILARLY,  $\Delta_{med}(1,3)$  is the median TWA amplitude of SCDDB population minus the median TWA AMPLITUDE IN THE NSRDB POPULATION. The median TWA AMPLITUDE IN THE NSRDB POPULATION. † INDICATES A SIGNIFICANT DIFFERENCE BETWEEN TWA AMPLITUDES AT A GIVEN HR RANGE USING THE KOLMOGOROV-SMIRNOV TEST (p < 0.0001). THE EMPTY ENTRIES (-) INDICATE THAT THERE WERE FEWER THAN 10 DETECTED EPISODES OF TWA ACTIVITY IN THE CORRESPONDING PATIENT POPULATIONS, AND THUS NOT AMENABLE TO SIGNIFICANCE TESTING USING THE KOLMOGOROV-SMIRNOV TEST.

HR Band	NSRDB/CHFDB/SCDDB	$\Delta_{med(1,2)}$	$\Delta_{med(1,3)}$
(beats/min)	(% indeterminate)	$(\mu V)$	$(\mu V)$
	SAM		
40-50	1±5 / - / 20±20	-	-2.28†
50-60	1±7 / 14±15 / 20±23	0.03†	-2.58†
60-70	4±11 / 7±12 / 16±20	-1.67†	-3.14†
70-80	4±11 / 8±14 / 10±16	-2.82†	-6.23†
80-90	10±17 / 14±18 / 8±16	-1.08†	-5.50†
90-100	14±19 / 11±18 / 13±20	3.05†	-5.70†
100-110	12±18 / 17±21 / 8±15	9.28†	-5.60†
110-120	$27\pm25$ / $14\pm14$ / $9\pm16$	25.05†	-1.35†
After Significance Testing (SDA method)			
40-50	88±4 / - / 80±11	-	-2.70†
50-60	86±9 / - / 88±8	-	18.41†
60-70	83±12 / 89±5 / 90±5	8.01†	6.47†
70-80	88±8 / 82±18 / 90±9	3.45†	0.11†
80-90	88±7 / 70±25 / 68±30	9.60†	6.95†
90-100	88±8 / 84±18 / 86±12	20.71†	0.43†
100-110	91±6 / 81±18 / 90±8	28.09†	9.69†
110-120	92±4 / 66±17 / 92±6	30.96†	41.51†

TABLE III

COMPARISON OF THE NSRDB, CHFDB, AND SCDDB USING THE SM WITH A SPECTRAL RATIO THRESHOLD VALUE OF k=3. See Table II for A description of the presented items.

HR Band (beats/min)	NSRDB/CHFDB/SCDDB (% indeterminate)	$\Delta_{med(1,2)} \atop (\mu \mathrm{V})$	$\Delta_{med(1,3)} \ (\mu \mathrm{V})$
40-50	- / - / -	-	-
50-60	$87{\pm}10$ / - / $92\pm6$	-	12.12†
60-70	$91{\pm}6$ / $85{\pm}15$ / $93{\pm}4$	-1.37†	-4.57†
70-80	93±4 / 78±21 / 91±6	3.91†	-5.79†
80-90	94±3 / 87±11 / 87±13	3.68†	-5.33
90-100	$95{\pm}3$ / $87{\pm}10$ / $88{\pm}9$	9.15†	-4.21†
100-110	$95{\pm}2$ / $76{\pm}13$ / $91{\pm}7$	28.14†	-4.25†
110-120	$95{\pm}2$ / $74{\pm}15$ / $90{\pm}8$	21.95†	3.24†

#### **IV. DISCUSSION**

TWA analysis generally leads to a large number of indeterminate cases [4], [6], [7]. Furthermore, 'natural' TWA activity of normal subjects of up to 10  $\mu$ V has been reported in healthy subjects [19]. The results of our study suggest that these observations may be explained by the high number of false positive TWA events, particularly during periods of higher noise (such as during exercise/stress test when the signal quality is qualitatively similar to the 10 dB simulated records studied here). In addition, our simulation study indicates that in the absence of appropriate (adaptive non-parametric) significance testing, even a relatively small amount of noise (due to muscle artifact, baseline wander or electrode motion) can lead to the raising of the noise floor to clinically significant levels (10  $\mu$ V or much more).

Our results on artificial data indicate that the SDA method produces a more accurate detection of TWA patterns in noise, when compared to other standard or more advanced techniques of noise rejection at both low and high values of TWA and

TABLE IV Comparison of the NSRDB, CHFDB, and SCDDB using the MMA method. See Table II for a description of the presented items.

HR Band	NSRDB/CHFDB/SCDDB	Δ (1.0)	Δ (1.0)
		$\rightarrow med(1,2)$	$\rightarrow med(1,3)$
(beats/min)	(% indeterminate)	$(\mu \mathbf{V})$	$(\mu \mathbf{V})$
40-50	1±5 / - / 20±20	-	-2.94†
50-60	$1\pm7$ / $14\pm15$ / $20\pm23$	-2.14	-2.63†
60-70	$4{\pm}11$ / $7{\pm}12$ / $16{\pm}20$	-1.56†	-4.28†
70-80	$4{\pm}11$ / $8{\pm}14$ / $10{\pm}16$	-4.54†	-8.64†
80-90	$10{\pm}17$ / $14{\pm}18$ / $8{\pm}16$	-2.33†	-7.85†
90-100	$14{\pm}19$ / $11{\pm}18$ / $13{\pm}20$	1.53†	-7.86†
100-110	$12{\pm}18$ / $17{\pm}21$ / $8{\pm}15$	9.36†	-7.92†
110-120	$27{\pm}25$ / $14{\pm}14$ / $9{\pm}16$	23.19†	-3.43†

noise. Since our technique assumes nothing concerning the noise distribution, we expect (and observe) a lower error rate. The inverse relationship between the false alarm rate and missed detection rate is well known; reducing one results in increasing the other and vice versa. Thus, to facilitate comparison of the three detection algorithms discussed in this work, we fixed their false alarm rates at the simulated TWA amplitude of  $0\mu V$  to approximately 1-2% and studied their missed detection rates. As summarized in Table I, the SDA method resulted in a statistically significant reduction in the percentage of missed detections (or indeterminate cases) at every simulated TWA amplitude from 50 to  $100\mu V$  (Wilcoxon rank sum test, p < 0.001).

The SAM method is utilized in this work as a baseline amplitude estimation technique to demonstrate the applicability of the SDA method for reducing false detections (false alarms) and simultaneously reducing missed detections. Our rational for choosing the SAM method was its ease of interpretation, as that it is based on simple averaging in time-domain. Logically, we can expect that more sophisticated TWA amplitude estimation techniques, in association with the SDA method, will result in further improvements.

When testing the effect of window size we found that decreasing the window length from 128 to 64 to 32 beats simply raised the noise floor, but did not affect the trend in our result. These observations are consistent with the previously reported results concerning the influence of window size on the performance of the SM and the MMA method [9]. Nevertheless, increasing the window size is not always practical, since one might wish to decrease the window length to mitigate the nonstationary effects such as phase changes due to ectopy and HR perturbations, and to be able to more rapidly track changes in TWA amplitude. We also demonstrate that the SM technique produces an estimate of the TWA amplitude that is biased towards values lower than simulated values. (This observation can be explained mathematically and is beyond the scope of this work [28], [32].)

Studies on both artificial data and three different patient populations (using the NSRDB, the CHFDB, and the SCCDB) indicate that our new detection algorithm provides enhanced discriminatory power between patient populations. (The median difference between healthy and unhealthy patients is significantly larger than the other standard techniques at almost all HRs (p < 0.0001).) The most marked differences are found at higher HRs, although HRs below standard thresholds

Copyright (c) 2010 IEEE. Personal use is permitted. For any other purposes, Permission must be obtained from the IEEE by emailing pubs-permissions@ieee.org. Authorized licensed use limited to: Universidad de Zaragoza. Downloaded on June 01,2010 at 10:43:47 UTC from IEEE Xplore. Restrictions apply. (110 BPM) also allow differentiation of normal and abnormal subjects. Note that in each case the application of significance testing increased the margin of separability between the patient populations (see Tables II and III). This improvement can be explained as follows: before significance testing the number of reported alternans-like episodes with relatively small amplitudes were much larger, and thus the quantiles were biased towards zero. After significance testing a large number of such episodes were marked as indeterminate and thus were removed from the quantile calculations, and therefore each quantile took on a larger value. Therefore, the application of significance testing improved the margin of separability by removing false detections that were negatively weighting the calculations. Furthermore, on average the surrogate test maintained a lower percentage of indeterminate case than a comparable test (the SM). This observation may be explained by the lower missed detection rate of our surrogate test method.

Note that the SM does include a test of significance (designed under assumption of Gaussianity of the spectral coefficients) while in contrast the MMA method relies on the preprocessing steps (such as exclusion of beats with abnormal fluctuations in the TP complex) to reduce noise artifacts [29], [30]. We repeated the TWA analysis on the real data using the SM and MMA methods. In the case of the NSRDB-CHFDB, application of the SM resulted in an improvement in inter-population separability over both the SAM (with no significance testing) and the MMA method. However, in the case of the NSRDB-SCDDB, only the SAM with the surrogate testing method was able to improve the inter-population margin of separability. The MMA method produced similar results to the SAM (with no significance testing) although with a positive offset at all HRs, as we would expect from our experience on artificial data.

It is worth noting that the SDA method is not computationally more expensive than the standard methods, since the bulk of the computation of TWA algorithms is devoted to preprocessing and alignment of beats, with generation of surrogate data through beat index reshuffling and re-estimation of TWA amplitude being only a small portion of the overall computational cost.

Finally, we note that we employed the same preprocessing for all the methods, using the best available open source algorithms [25] since our work focuses on significance testing. The commercial implementations of the MMA and the SM may include additional or alternative pre-processing and noise reduction steps that are not considered here. However, it is unlikely that even extremely sophisticated preprocessing or estimation methods (such as complex demodulation or timefrequency approaches) would obviate the need for significance testing since there exits no known technique which completely removes all noise in the ECG.

#### V. CONCLUSIONS

We have described a new application of a non-parametric surrogate test to reject false TWA-like activity (which could have been due to artifacts or noise). The new technique was evaluated on both real and artificial data. Tests on the artificial data demonstrate the superiority of our method over existing TWA detection methods both at low and high levels of TWA amplitude.

In the absence of background physiological noise, a lower noise floor of 5-10  $\mu$ V was found, below which the measured TWA is unreliable and could be due to noise alone. This noise floor may account for some reports of TWA in normal patients below 10  $\mu$ V. Results also demonstrate that the higher the background noise, the more likely it is that a given technique will falsely detect TWA and over-estimate the magnitude of the TWA.

When evaluated on three public databases (the NSRDB, CHFDB and SCDDB) our new approach demonstrated significant differences in TWA amplitudes between each database at all HRs intervals between 40 BPM and 120 BPM. The most marked differences were generally found at higher HRs, and the new technique provided a larger margin of separability between patient populations than the standard methods. Our results also indicate that population separation is possible at lower HRs than currently clinically recommended.

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